

# Singularity of random Bernoulli 0/1 matrices

**Alexander Litvak**

University of Alberta

based on a work in progress with

**K. Tikhomirov**

EIMI, St. Petersburg, 2019

# Random $\pm 1$ matrices

**An old problem:** Let  $B$  be an  $n \times n$  random matrix with i.i.d.  $\pm 1$  entries. What is

$$P_n := \mathbb{P}\{B \text{ is singular}\}?$$

# Random $\pm 1$ matrices

**An old problem:** Let  $B$  be an  $n \times n$  random matrix with i.i.d.  $\pm 1$  entries. What is

$$P_n := \mathbb{P}\{B \text{ is singular}\}?$$

**Equivalently:** Let  $X_1, X_2, \dots, X_n$  be independent random vectors uniformly distributed on the vertices of the  $n$ -dimensional cube  $[-1, 1]^n$ .

# Random $\pm 1$ matrices

**An old problem:** Let  $B$  be an  $n \times n$  random matrix with i.i.d.  $\pm 1$  entries. What is

$$P_n := \mathbb{P}\{B \text{ is singular}\}?$$

**Equivalently:** Let  $X_1, X_2, \dots, X_n$  be independent random vectors uniformly distributed on the vertices of the  $n$ -dimensional cube  $[-1, 1]^n$ .

What is the probability that the vectors are linearly dependent?

# Random $\pm 1$ matrices

**An old problem:** Let  $B$  be an  $n \times n$  random matrix with i.i.d.  $\pm 1$  entries. What is

$$P_n := \mathbb{P}\{B \text{ is singular}\}?$$

**Equivalently:** Let  $X_1, X_2, \dots, X_n$  be independent random vectors uniformly distributed on the vertices of the  $n$ -dimensional cube  $[-1, 1]^n$ .

What is the probability that the vectors are linearly dependent?

**The trivial lower bound:**

$$P_n \geq \mathbb{P}\{\text{Two rows/columns of } B \text{ are equal up to a sign}\} \geq (1 - o(1)) 2n^2 2^{-n}.$$

# Random $\pm 1$ matrices

**An old problem:** Let  $B$  be an  $n \times n$  random matrix with i.i.d.  $\pm 1$  entries. What is

$$P_n := \mathbb{P}\{B \text{ is singular}\}?$$

**Equivalently:** Let  $X_1, X_2, \dots, X_n$  be independent random vectors uniformly distributed on the vertices of the  $n$ -dimensional cube  $[-1, 1]^n$ .

What is the probability that the vectors are linearly dependent?

**The trivial lower bound:**

$$P_n \geq \mathbb{P}\{\text{Two rows/columns of } B \text{ are equal up to a sign}\} \geq (1 - o(1)) 2n^2 2^{-n}.$$

**A natural conjecture:** This is the main reason for singularity.

# Random $\pm 1$ matrices

**An old problem:** Let  $B$  be an  $n \times n$  random matrix with i.i.d.  $\pm 1$  entries. What is

$$P_n := \mathbb{P}\{B \text{ is singular}\}?$$

**Equivalently:** Let  $X_1, X_2, \dots, X_n$  be independent random vectors uniformly distributed on the vertices of the  $n$ -dimensional cube  $[-1, 1]^n$ .

What is the probability that the vectors are linearly dependent?

**The trivial lower bound:**

$$P_n \geq \mathbb{P}\{\text{Two rows/columns of } B \text{ are equal up to a sign}\} \geq (1 - o(1)) 2n^2 2^{-n}.$$

**A natural conjecture:** This is the main reason for singularity.

**Conjecture 1:**  $P_n = (1/2 + o(1))^n = 2^{-(1+o(1))n}.$

# Random $\pm 1$ matrices

**An old problem:** Let  $B$  be an  $n \times n$  random matrix with i.i.d.  $\pm 1$  entries. What is

$$P_n := \mathbb{P}\{B \text{ is singular}\}?$$

**Equivalently:** Let  $X_1, X_2, \dots, X_n$  be independent random vectors uniformly distributed on the vertices of the  $n$ -dimensional cube  $[-1, 1]^n$ .

What is the probability that the vectors are linearly dependent?

**The trivial lower bound:**

$$P_n \geq \mathbb{P}\{\text{Two rows/columns of } B \text{ are equal up to a sign}\} \geq (1 - o(1)) 2n^2 2^{-n}.$$

**A natural conjecture:** This is the main reason for singularity.

**Conjecture 1:** 
$$P_n = (1/2 + o(1))^n = 2^{-(1+o(1))n}.$$

**Conjecture 2:** 
$$P_n = (1+o(1)) 2n^2 2^{-n}.$$



Komlós (67):  $P_n \rightarrow 0$ .

# Known results

Komlós (67):  $P_n \rightarrow 0$ .

Kahn, Komlós and Szemerédi (95):  $P_n \leq 0.999^n$ .

# Known results

Komlós (67):  $P_n \rightarrow 0$ .

Kahn, Komlós and Szemerédi (95):  $P_n \leq 0.999^n$ .

Tao–Vu (07):  $P_n \leq (3/4 + o(1))^n$ .

# Known results

Komlós (67):  $P_n \rightarrow 0$ .

Kahn, Komlós and Szemerédi (95):  $P_n \leq 0.999^n$ .

Tao–Vu (07):  $P_n \leq (3/4 + o(1))^n$ .

Bourgain–Vu–P.M. Wood (10):  $P_n \leq (1/\sqrt{2} + o(1))^n$ .

# Known results

Komlós (67):  $P_n \rightarrow 0$ .

Kahn, Komlós and Szemerédi (95):  $P_n \leq 0.999^n$ .

Tao–Vu (07):  $P_n \leq (3/4 + o(1))^n$ .

Bourgain–Vu–P.M. Wood (10):  $P_n \leq (1/\sqrt{2} + o(1))^n$ .

K. Tikhomirov (19+):  $P_n \leq (1/2 + o(1))^n$ , solving Conjecture 1.

# Bernoulli random matrices

One can ask a similar question about Bernoulli 0/1 random matrices:

# Bernoulli random matrices

One can ask a similar question about Bernoulli 0/1 random matrices:

Let  $p \in (0, 1/2]$  and let  $B_p$  be an  $n \times n$  random matrix with i.i.d. 0/1 random variables taking value 1 with probability  $p$ .

# Bernoulli random matrices

One can ask a similar question about Bernoulli 0/1 random matrices:

Let  $p \in (0, 1/2]$  and let  $B_p$  be an  $n \times n$  random matrix with i.i.d. 0/1 random variables taking value 1 with probability  $p$ . Note that  $B_p$  can be viewed as the adjacency matrix of [Erdős–Rényi](#) graph — a random graph on  $n$  vertices whose edges appear independently of others with probability  $p$ .



# Bernoulli random matrices

One can ask a similar question about Bernoulli 0/1 random matrices:

Let  $p \in (0, 1/2]$  and let  $B_p$  be an  $n \times n$  random matrix with i.i.d. 0/1 random variables taking value 1 with probability  $p$ . Note that  $B_p$  can be viewed as the adjacency matrix of [Erdős–Rényi](#) graph — a random graph on  $n$  vertices whose edges appear independently of others with probability  $p$ .

**Question:** What is

$$P_n := \mathbb{P} \{B_p \text{ is singular}\}?$$

# Bernoulli random matrices

One can ask a similar question about Bernoulli 0/1 random matrices:

Let  $p \in (0, 1/2]$  and let  $B_p$  be an  $n \times n$  random matrix with i.i.d. 0/1 random variables taking value 1 with probability  $p$ . Note that  $B_p$  can be viewed as the adjacency matrix of [Erdős–Rényi](#) graph — a random graph on  $n$  vertices whose edges appear independently of others with probability  $p$ .

**Question:** What is

$$P_n := \mathbb{P} \{B_p \text{ is singular}\}?$$

**Conjecture:**

$$P_n = (1 + o(1)) \mathbb{P} \{ \exists \text{ a zero row or a zero column} \} = (1 + o(1)) 2n(1 - p)^n.$$

**Geometrically** the condition means that either  $\exists$  a zero column or  $\exists$  a *coordinate* hyperplane such that all columns belong to it.

# Bernoulli random matrices

One can ask a similar question about Bernoulli 0/1 random matrices:

Let  $p \in (0, 1/2]$  and let  $B_p$  be an  $n \times n$  random matrix with i.i.d. 0/1 random variables taking value 1 with probability  $p$ . Note that  $B_p$  can be viewed as the adjacency matrix of [Erdős–Rényi](#) graph — a random graph on  $n$  vertices whose edges appear independently of others with probability  $p$ .

**Question:** What is

$$P_n := \mathbb{P} \{B_p \text{ is singular}\}?$$

**Conjecture:**

$$P_n = (1 + o(1)) \mathbb{P} \{ \exists \text{ a zero row or a zero column} \} = (1 + o(1)) 2n(1 - p)^n.$$

**Geometrically** the condition means that either  $\exists$  a zero column or  $\exists$  a *coordinate* hyperplane such that all columns belong to it.

Many works on different models of sparse matrices (with iid entries):

[Götze–A. Tikhomirov](#), [Costello–Vu](#), [Basak–Rudelson](#), [Rudelson–K. Tikhomirov](#),  
[Tao–Vu](#).

# Bernoulli random matrices

Basak–Rudelson (17):  $P_n \leq \exp(-cnp)$  for  $p = p(n) \geq (C \ln n)/n$ ,

# Bernoulli random matrices

**Basak–Rudelson (17):**  $P_n \leq \exp(-cnp)$  for  $p = p(n) \geq (C \ln n)/n$ , moreover

$$\mathbb{P} \left\{ s_n(B_p) \leq c \exp(-C \ln(1/p) / \ln(np)) t \sqrt{p/n} \right\} \leq t + \exp(-cnp),$$

where

$$s_n(M) = \inf_{|x|=1} |Mx|.$$

# Bernoulli random matrices

**Basak–Rudelson (17):**  $P_n \leq \exp(-cnp)$  for  $p = p(n) \geq (C \ln n)/n$ , moreover

$$\mathbb{P} \left\{ s_n(B_p) \leq c \exp(-C \ln(1/p) / \ln(np)) t \sqrt{p/n} \right\} \leq t + \exp(-cnp),$$

where

$$s_n(M) = \inf_{|x|=1} |Mx|.$$

**K. Tikhomirov (19+):**  $P_n \leq (1 - p + o(1))^n$  for  $p \in (0, 1/2]$  (independent of  $n$ ).

# Bernoulli random matrices

**Basak–Rudelson (17):**  $P_n \leq \exp(-cnp)$  for  $p = p(n) \geq (C \ln n)/n$ , moreover

$$\mathbb{P} \left\{ s_n(B_p) \leq c \exp(-C \ln(1/p) / \ln(np)) t \sqrt{p/n} \right\} \leq t + \exp(-cnp),$$

where

$$s_n(M) = \inf_{|x|=1} |Mx|.$$

**K. Tikhomirov (19+):**  $P_n \leq (1 - p + o(1))^n$  for  $p \in (0, 1/2]$  (independent of  $n$ ).

Moreover,  $\forall \varepsilon > 0, \forall n \geq n(p, \varepsilon)$ ,

$$\mathbb{P} \left\{ s_n(B_p) \leq t \sqrt{p/n} \right\} \leq C(p, \varepsilon)t + (1 - p + \varepsilon)^n.$$

## L-K.T.

There is a (small) absolute constant  $c > 0$  such that the following holds.  
Let  $0 < q < c$  be a parameter and  $q < p < c$ . Then,

$$P_n \leq (1 + o_q(1)) 2n(1 - p)^n.$$



# Main result

## L-K.T.

There is a (small) absolute constant  $c > 0$  such that the following holds.  
Let  $0 < q < c$  be a parameter and  $q < p < c$ . Then,

$$P_n \leq (1 + o_q(1)) 2n(1 - p)^n.$$

Moreover,  $\forall \varepsilon > 0, \forall n \geq n(p, \varepsilon)$ ,

$$\mathbb{P} \left\{ s_n(B_p) \leq t n^{-c_q} \right\} \leq t + (1 + o_q(1)) 2n(1 - p)^n.$$

# Main result

## L-K.T.

There is a (small) absolute constant  $c > 0$  such that the following holds. Let  $0 < q < c$  be a parameter and  $q < p < c$ . Then,

$$P_n \leq (1 + o_q(1)) 2n(1 - p)^n.$$

Moreover,  $\forall \varepsilon > 0, \forall n \geq n(p, \varepsilon)$ ,

$$\mathbb{P} \{s_n(B_p) \leq t n^{-C_q}\} \leq t + (1 + o_q(1)) 2n(1 - p)^n.$$

**Remark.** It seems that using technique from [LLTTY](#) ([Lytova](#), [Tomczak-Jaegermann](#), [Youssef + LT](#)) papers on random regular matrices one can substitute the assumption  $q < p < c$  with  $\frac{C \ln n}{n} < p < c$  and to remove dependence of the constants on  $q$ . (A 0/1 matrix is regular if the sums of 1 in all columns and in all rows are the same — it is the adjacency matrix of a regular directed graph).

# Some ideas of the proof.

It is well-understood by now that to deal with the smallest singular number one needs to split  $S^{n-1}$  into several parts and to work separately on each part.

# Some ideas of the proof.

It is well-understood by now that to deal with the smallest singular number one needs to split  $S^{n-1}$  into several parts and to work separately on each part.

This idea goes back to [Kashin 77](#), where, in order to obtain an orthogonal decomposition of  $\ell_1^n$ , he split the sphere into two classes according to the ratio of  $\ell_1^n$  and  $\ell_2^n$  norms. In a similar context it was used by [Schehtman 04](#).

# Some ideas of the proof.

It is well-understood by now that to deal with the smallest singular number one needs to split  $S^{n-1}$  into several parts and to work separately on each part.

This idea goes back to [Kashin 77](#), where, in order to obtain an orthogonal decomposition of  $\ell_1^n$ , he split the sphere into two classes according to the ratio of  $\ell_1^n$  and  $\ell_2^n$  norms. In a similar context it was used by [Schehtman 04](#).

Since we want to provide a lower bound on the smallest singular value of a random matrix  $M$ , we need to show that  $|Mx|$  is not very small for all  $x \in S^{n-1}$ . Usually it is done using the union bound – to prove a good probability bound for an individual vector  $x$  and then to find a good net in order to apply approximation. The main point is to have a good balance between the probability and the cardinality of a net.

# Some ideas of the proof.

But vectors behave differently. Consider the following example, let  $X = \{\varepsilon_i\}$  be a Bernoulli random vector with  $\pm 1$  independent entries. Then

$$\langle X, e_1 + e_2 \rangle = \varepsilon_1 + \varepsilon_2 = 0 \quad \text{with probability} \quad 1/2.$$

On the other hand,

$$\langle X, \sum_i e_i \rangle = \sum_i \varepsilon_i = 0 \quad \text{with probability at most} \quad 1/\sqrt{n}$$

by the [Erdős–Littlewood–Offord](#) anti-concentration lemma.

# Some ideas of the proof.

But vectors behave differently. Consider the following example, let  $X = \{\varepsilon_i\}$  be a Bernoulli random vector with  $\pm 1$  independent entries. Then

$$\langle X, e_1 + e_2 \rangle = \varepsilon_1 + \varepsilon_2 = 0 \quad \text{with probability} \quad 1/2.$$

On the other hand,

$$\langle X, \sum_i e_i \rangle = \sum_i \varepsilon_i = 0 \quad \text{with probability at most} \quad 1/\sqrt{n}$$

by the [Erdős–Littlewood–Offord](#) anti-concentration lemma.

Usually, it is hard to get good individual bounds for vectors of small support, so-called *sparse vectors*.

# Some ideas of the proof.

But vectors behave differently. Consider the following example, let  $X = \{\varepsilon_i\}$  be a Bernoulli random vector with  $\pm 1$  independent entries. Then

$$\langle X, e_1 + e_2 \rangle = \varepsilon_1 + \varepsilon_2 = 0 \quad \text{with probability} \quad 1/2.$$

On the other hand,

$$\langle X, \sum_i e_i \rangle = \sum_i \varepsilon_i = 0 \quad \text{with probability at most} \quad 1/\sqrt{n}$$

by the [Erdős–Littlewood–Offord](#) anti-concentration lemma.

Usually, it is hard to get good individual bounds for vectors of small support, so-called *sparse vectors*. However, the set of such vectors is essentially of lower dimension, hence admit a very good net.



# Some ideas of the proof.

But vectors behave differently. Consider the following example, let  $X = \{\varepsilon_i\}$  be a Bernoulli random vector with  $\pm 1$  independent entries. Then

$$\langle X, e_1 + e_2 \rangle = \varepsilon_1 + \varepsilon_2 = 0 \quad \text{with probability} \quad 1/2.$$

On the other hand,

$$\langle X, \sum_i e_i \rangle = \sum_i \varepsilon_i = 0 \quad \text{with probability at most} \quad 1/\sqrt{n}$$

by the [Erdős–Littlewood–Offord](#) anti-concentration lemma.

Usually, it is hard to get good individual bounds for vectors of small support, so-called *sparse vectors*. However, the set of such vectors is essentially of lower dimension, hence admit a very good net. This leads to splitting the sphere into *compressible* vectors – those closed to sparse, and *incompressible* vectors – the rest.

# Some ideas of the proof.

But vectors behave differently. Consider the following example, let  $X = \{\varepsilon_i\}$  be a Bernoulli random vector with  $\pm 1$  independent entries. Then

$$\langle X, e_1 + e_2 \rangle = \varepsilon_1 + \varepsilon_2 = 0 \quad \text{with probability} \quad 1/2.$$

On the other hand,

$$\langle X, \sum_i e_i \rangle = \sum_i \varepsilon_i = 0 \quad \text{with probability at most} \quad 1/\sqrt{n}$$

by the [Erdős–Littlewood–Offord](#) anti-concentration lemma.

Usually, it is hard to get good individual bounds for vectors of small support, so-called *sparse vectors*. However, the set of such vectors is essentially of lower dimension, hence admit a very good net. This leads to splitting the sphere into *compressible* vectors – those closed to sparse, and *incompressible* vectors – the rest. For compressible vectors we have a net of small cardinality, therefore relatively poor individual probability bounds work, while incompressible vectors are well spread and therefore have very good anti-concentration properties.

# Some ideas of the proof.

But vectors behave differently. Consider the following example, let  $X = \{\varepsilon_i\}$  be a Bernoulli random vector with  $\pm 1$  independent entries. Then

$$\langle X, e_1 + e_2 \rangle = \varepsilon_1 + \varepsilon_2 = 0 \quad \text{with probability} \quad 1/2.$$

On the other hand,

$$\langle X, \sum_i e_i \rangle = \sum_i \varepsilon_i = 0 \quad \text{with probability at most} \quad 1/\sqrt{n}$$

by the [Erdős–Littlewood–Offord](#) anti-concentration lemma.

Usually, it is hard to get good individual bounds for vectors of small support, so-called *sparse vectors*. However, the set of such vectors is essentially of lower dimension, hence admit a very good net. This leads to splitting the sphere into *compressible* vectors – those closed to sparse, and *incompressible* vectors – the rest. For compressible vectors we have a net of small cardinality, therefore relatively poor individual probability bounds work, while incompressible vectors are well spread and therefore have very good anti-concentration properties. This approach was used in [L–Pajor–Rudelson–Tomczak-Jaegermann \(05\)](#) for rectangular matrices and was later developed in series of works by [Rudelson–Vershynin](#).

# Some ideas of the proof.

For 0/1 matrices an additional problem is caused by constant vectors. Indeed, while properly normalized centered random matrices (say with entries  $\pm 1$ ) have norm of order  $\sqrt{n}$ , the norm  $\|B_p\| \approx pn$ .

# Some ideas of the proof.

For 0/1 matrices an additional problem is caused by constant vectors. Indeed, while properly normalized centered random matrices (say with entries  $\pm 1$ ) have norm of order  $\sqrt{n}$ , the norm  $\|B_p\| \approx pn$ . Fortunately, this large norm is only in the direction of  $\mathbf{1} = (1, 1, \dots, 1)$ . On the subspace orthogonal to  $\mathbf{1}$  the norm is of the order  $\sqrt{pn}$ .

# Some ideas of the proof.

For 0/1 matrices an additional problem is caused by constant vectors. Indeed, while properly normalized centered random matrices (say with entries  $\pm 1$ ) have norm of order  $\sqrt{n}$ , the norm  $\|B_p\| \approx pn$ . Fortunately, this large norm is only in the direction of  $\mathbf{1} = (1, 1, \dots, 1)$ . On the subspace orthogonal to  $\mathbf{1}$  the norm is of the order  $\sqrt{pn}$ .

This leads to our splitting. The first class will be sparse vectors shifted by constants vectors. The second class will be the remaining vectors.

# Some ideas of the proof.

For 0/1 matrices an additional problem is caused by constant vectors. Indeed, while properly normalized centered random matrices (say with entries  $\pm 1$ ) have norm of order  $\sqrt{n}$ , the norm  $\|B_p\| \approx pn$ . Fortunately, this large norm is only in the direction of  $\mathbf{1} = (1, 1, \dots, 1)$ . On the subspace orthogonal to  $\mathbf{1}$  the norm is of the order  $\sqrt{pn}$ .

This leads to our splitting. The first class will be sparse vectors shifted by constants vectors. The second class will be the remaining vectors.

For the first class standard anti-concentration technique works, since the set is essentially of lower dimension (although there are many cases).

# Some ideas of the proof.

For the second class we show that it is contained in *gradual non-constant vectors*, that is, vectors (after certain normalization and for some parameters  $r, \delta, L, h$ ) s.t.

1.  $x_{rn}^* = 1$
2.  $x_i^* \leq (n/i)^L$
3. If  $(y_i)_i$  is a non-increasing rearrangement of  $(x_i)_i$  then  $y_{\delta n} - y_{n-\delta n} \geq h$ .



# Some ideas of the proof.

For the second class we show that it is contained in *gradual non-constant vectors*, that is, vectors (after certain normalization and for some parameters  $r, \delta, L, h$ ) s.t.

1.  $x_{rn}^* = 1$
2.  $x_i^* \leq (n/i)^L$
3. If  $(y_i)_i$  is a non-increasing rearrangement of  $(x_i)_i$  then  $y_{\delta n} - y_{n-\delta n} \geq h$ .

To work with this class we partially follow [Rudelson–Vershynin](#) scheme.

# Some ideas of the proof.

For the second class we show that it is contained in *gradual non-constant vectors*, that is, vectors (after certain normalization and for some parameters  $r, \delta, L, h$ ) s.t.

1.  $x_{rn}^* = 1$
2.  $x_i^* \leq (n/i)^L$
3. If  $(y_i)_i$  is a non-increasing rearrangement of  $(x_i)_i$  then  $y_{\delta n} - y_{n-\delta n} \geq h$ .

To work with this class we partially follow [Rudelson–Vershynin](#) scheme.

First, one reduces estimating the smallest singular number to estimating distances between a column  $X_i$  to the span of remaining columns, say  $H_i$ ,  $i \leq 1$ .

# Some ideas of the proof.

For the second class we show that it is contained in *gradual non-constant vectors*, that is, vectors (after certain normalization and for some parameters  $r, \delta, L, h$ ) s.t.

1.  $x_{rn}^* = 1$
2.  $x_i^* \leq (n/i)^L$
3. If  $(y_i)_i$  is a non-increasing rearrangement of  $(x_i)_i$  then  $y_{\delta n} - y_{n-\delta n} \geq h$ .

To work with this class we partially follow [Rudelson–Vershynin](#) scheme.

First, one reduces estimating the smallest singular number to estimating distances between a column  $X_i$  to the span of remaining columns, say  $H_i$ ,  $i \leq 1$ .

This distance is a projection on a (random) normal vector to  $H_i$ .

# Some ideas of the proof.

For the second class we show that it is contained in *gradual non-constant vectors*, that is, vectors (after certain normalization and for some parameters  $r, \delta, L, h$ ) s.t.

1.  $x_{rn}^* = 1$
2.  $x_i^* \leq (n/i)^L$
3. If  $(y_i)_i$  is a non-increasing rearrangement of  $(x_i)_i$  then  $y_{\delta n} - y_{n-\delta n} \geq h$ .

To work with this class we partially follow [Rudelson–Vershynin](#) scheme.

First, one reduces estimating the smallest singular number to estimating distances between a column  $X_i$  to the span of remaining columns, say  $H_i$ ,  $i \leq 1$ .

This distance is a projection on a (random) normal vector to  $H_i$ .

Thus, we have an inner product of  $X_i$  and the normal (note that they are independent).

# Some ideas of the proof.

For the second class we show that it is contained in *gradual non-constant vectors*, that is, vectors (after certain normalization and for some parameters  $r, \delta, L, h$ ) s.t.

1.  $x_{rn}^* = 1$
2.  $x_i^* \leq (n/i)^L$
3. If  $(y_i)_i$  is a non-increasing rearrangement of  $(x_i)_i$  then  $y_{\delta n} - y_{n-\delta n} \geq h$ .

To work with this class we partially follow [Rudelson–Vershynin](#) scheme.

First, one reduces estimating the smallest singular number to estimating distances between a column  $X_i$  to the span of remaining columns, say  $H_i$ ,  $i \leq 1$ .

This distance is a projection on a (random) normal vector to  $H_i$ .

Thus, we have an inner product of  $X_i$  and the normal (note that they are independent).

Then we apply an anti-concentration property (such a property says that an inner product of a random vector with a flat vector can't concentrate around a number).

# Some ideas of the proof.

For the second class we show that it is contained in *gradual non-constant vectors*, that is, vectors (after certain normalization and for some parameters  $r, \delta, L, h$ ) s.t.

1.  $x_{rn}^* = 1$
2.  $x_i^* \leq (n/i)^L$
3. If  $(y_i)_i$  is a non-increasing rearrangement of  $(x_i)_i$  then  $y_{\delta n} - y_{n-\delta n} \geq h$ .

To work with this class we partially follow [Rudelson–Vershynin](#) scheme.

First, one reduces estimating the smallest singular number to estimating distances between a column  $X_i$  to the span of remaining columns, say  $H_i$ ,  $i \leq 1$ .

This distance is a projection on a (random) normal vector to  $H_i$ .

Thus, we have an inner product of  $X_i$  and the normal (note that they are independent).

Then we apply an anti-concentration property (such a property says that an inner product of a random vector with a flat vector can't concentrate around a number).

To make this scheme work, [Rudelson–Vershynin](#) introduced LCD (*least common denominator*), which, in a sense, measures how close a proportional coordinate projection of a vector to the properly rescaled integer lattice. They also had to develop [Littlewood–Offord](#) theory.

# Some ideas of the proof.

In our case both, the LCD, and the known anti-concentration results are not strong enough, so we need to develop new tools.

# Some ideas of the proof.

In our case both, the LCD, and the known anti-concentration results are not strong enough, so we need to develop new tools.

First idea is to pass from a Bernoulli random vector, which may have many zeros, to a random 0/1 vector with prescribed number of ones, say, with  $m$  ones, where  $m$  is of the order  $pn$ . Note that  $pn$  is an average number of ones in a Bernoulli vector.



# Some ideas of the proof.

In our case both, the LCD, and the known anti-concentration results are not strong enough, so we need to develop new tools.

First idea is to pass from a Bernoulli random vector, which may have many zeros, to a random 0/1 vector with prescribed number of ones, say, with  $m$  ones, where  $m$  is of the order  $pn$ . Note that  $pn$  is an average number of ones in a Bernoulli vector.

Second idea is to substitute LCD with another parameter, which we call balancing degree of a vector, and which is more directly related to the [Esseen](#) lemma, used to prove an anti-concentration.

# Some ideas of the proof.

In our case both, the LCD, and the known anti-concentration results are not strong enough, so we need to develop new tools.

First idea is to pass from a Bernoulli random vector, which may have many zeros, to a random 0/1 vector with prescribed number of ones, say, with  $m$  ones, where  $m$  is of the order  $pn$ . Note that  $pn$  is an average number of ones in a Bernoulli vector.

Second idea is to substitute LCD with another parameter, which we call balancing degree of a vector, and which is more directly related to the [Esseen](#) lemma, used to prove an anti-concentration.

Next we have to prove a [Littlewood–Offord](#) type anti-concentration property for this new parameter.

# Some ideas of the proof.

In our case both, the LCD, and the known anti-concentration results are not strong enough, so we need to develop new tools.

First idea is to pass from a Bernoulli random vector, which may have many zeros, to a random 0/1 vector with prescribed number of ones, say, with  $m$  ones, where  $m$  is of the order  $pn$ . Note that  $pn$  is an average number of ones in a Bernoulli vector.

Second idea is to substitute LCD with another parameter, which we call balancing degree of a vector, and which is more directly related to the [Esseen](#) lemma, used to prove an anti-concentration.

Next we have to prove a [Littlewood–Offord](#) type anti-concentration property for this new parameter.

In particular, we also extend the [Littlewood–Offord](#) theory to the case of dependent random variables (in our case – the coordinates of a vector with fixed number of ones).

# Balancing degree

Recall the definition of **Lévy** concentration function:

$$\mathcal{Q}(\xi, t) = \max_{\lambda} \mathbb{P}(|\xi - \lambda| < t).$$

# Balancing degree

Recall the definition of **Lévy** concentration function:

$$\mathcal{Q}(\xi, t) = \max_{\lambda} \mathbb{P}(|\xi - \lambda| < t).$$

**Esseen** Lemma (**66**):

$$\mathcal{Q}\left(\sum_{i=1}^m \xi_i, \tau\right) \leq C' \int_{-1}^1 \prod_{i=1}^m |\mathbb{E} \exp(2\pi \mathbf{i} \xi_i s / \tau)| ds.$$

# Balancing degree

Recall the definition of **Lévy** concentration function:

$$\mathcal{Q}(\xi, t) = \max_{\lambda} \mathbb{P}(|\xi - \lambda| < t).$$

**Esseen** Lemma (**66**):

$$\mathcal{Q}\left(\sum_{i=1}^m \xi_i, \tau\right) \leq C' \int_{-1}^1 \prod_{i=1}^m |\mathbb{E} \exp(2\pi \mathbf{i} \xi_i s / \tau)| ds.$$

For a finite integer subset  $S$ , let  $\eta[S]$  denotes a r.v. uniformly distributed on  $S$ .

# Balancing degree

Recall the definition of **Lévy** concentration function:

$$\mathcal{Q}(\xi, t) = \max_{\lambda} \mathbb{P}(|\xi - \lambda| < t).$$

**Esseen** Lemma (66):

$$\mathcal{Q}\left(\sum_{i=1}^m \xi_i, \tau\right) \leq C' \int_{-1}^1 \prod_{i=1}^m |\mathbb{E} \exp(2\pi i \xi_i s / \tau)| ds.$$

For a finite integer subset  $S$ , let  $\eta[S]$  denotes a r.v. uniformly distributed on  $S$ . Then

$$\mathbf{Bal}_n(v, m, K) := \sup \left\{ t > 0 : \frac{1}{N} \sum_{(S_1, \dots, S_m)} \int_{-t}^t \prod_{i=1}^m |\mathbb{E} \exp(2\pi i v_{\eta[S_i]} m^{-1/2} s)| ds \leq K \right\},$$

where the sum is taken over all sequences  $(S_i)_{i=1}^m$  of disjoint subsets  $S_1, \dots, S_m \subset [n]$ , each of cardinality  $\lfloor n/m \rfloor$ ,  $N$  is the number of such sequences,  $K \geq 1$  is a parameter.

# Balancing degree

Recall the definition of **Lévy** concentration function:

$$\mathcal{Q}(\xi, t) = \max_{\lambda} \mathbb{P}(|\xi - \lambda| < t).$$

**Esseen** Lemma (66):

$$\mathcal{Q}\left(\sum_{i=1}^m \xi_i, \tau\right) \leq C' \int_{-1}^1 \prod_{i=1}^m |\mathbb{E} \exp(2\pi i \xi_i s / \tau)| ds.$$

For a finite integer subset  $S$ , let  $\eta[S]$  denotes a r.v. uniformly distributed on  $S$ . Then

$$\mathbf{Bal}_n(v, m, K) := \sup \left\{ t > 0 : \frac{1}{N} \sum_{(S_1, \dots, S_m)} \int_{-t}^t \prod_{i=1}^m |\mathbb{E} \exp(2\pi i v_{\eta[S_i]} m^{-1/2} s)| ds \leq K \right\},$$

where the sum is taken over all sequences  $(S_i)_{i=1}^m$  of disjoint subsets  $S_1, \dots, S_m \subset [n]$ , each of cardinality  $\lfloor n/m \rfloor$ ,  $N$  is the number of such sequences,  $K \geq 1$  is a parameter.

We prove that

$$\mathcal{Q}\left(\sum_{i=1}^n v_i X_i, \sqrt{m} t\right) \leq C (t + 1/\mathbf{Bal}_n(v, m, K)) \quad \text{for all } t > 0,$$